Diagnosing hyperuniformity in two-dimensional, disordered, jammed packings of soft spheres

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I. INTRODUCTION

Hyperuniformity is a state of matter characterized by density fluctuations (when appropriately scaled) that vanish at large length scales [1], and the concept of hyperuniformity has emerged as a new way to classify crystals, quasicrystals, and special disordered systems [1,2]. Disordered hyperuniform materials, for example, behave more like crystals in the manner they suppress density fluctuations over large length scales, and yet they also resemble traditional liquids and glasses with statistically isotropic structures with no Bragg peaks. During the last decade, hyperuniform disordered states have been identified in maximally random jammed packings of hard particles [3–5], jammed athermal granular systems [6], jammed thermal colloidal packings [7], cold atoms [8], certain Coulombic systems [9], and “stealthy” disordered classical ground states [10]. Furthermore, the hyperuniformity property has been suggested to endow materials with novel physical properties potentially important for applications in photonics [11–13] and electronics [14–16]. Thus, the ability to ascertain the degree of hyperuniformity in a disordered system is becoming increasingly important.

Colloids offer a potentially fruitful experimental system for design, fabrication, and testing of hyperuniform systems. In fact, it has been conjectured that certain jammed particle packings are in the hyperuniform state [1]. If true, a significant laboratory challenge would be to learn how one might self-assemble jammed colloids into “stealthy” [10] hyperuniform states. The task of ascertaining whether or not a colloidal packing is hyperuniform is challenging as demonstrated by Berthier et al. [6]. This is due to the fact that typical experimental data sets have imperfections due to optical imaging and particle-size polydispersity. Thus, we currently lack a general method to test for hyperuniformity in experimental colloidal packings recorded via standard real-space methods such as optical microscopy.

To this end, we employ video optical microscopy to study hyperuniformity phenomena in disordered two-dimensional (2D) disordered, jammed packings of soft spheres. Using a combination of experiment and numerical simulation we characterize the detrimental effects of particle polydispersity, image noise, and finite-size effects on the assignment of hyperuniformity, and we develop a methodology that permits improved diagnosis of hyperuniformity from real-space measurements. The key to this improvement is a simple packing reconstruction algorithm that incorporates particle polydispersity to minimize the free volume. In addition, simulations show that hyperuniformity in finite-sized samples can be ascertained more accurately in direct space than in reciprocal space. Finally, our experimental colloidal packings of soft polymeric spheres are shown to be effectively hyperuniform.
II. HYPERUNIFORMITY

A. Point configurations

The hyperuniformity concept applies to many-particle systems with number density $\rho$ in $d$-dimensional Euclidean space $\mathbb{R}^d$ and is intimately related to the suppression of local density fluctuations at long length scales. In order to understand the concept more precisely, consider random placements of a spherical observation window $\Omega$ of radius $R$ (e.g., a circle in $d = 2$ or a sphere in $d = 3$) in a point pattern (e.g., a lattice of points or centers of the atoms of a liquid). The number of points, $N(R)$, contained in $\Omega$ is a random variable, and we define the number variance $\sigma_N^2(R)$ in an observation window of radius $R$ to be

$$\sigma_N^2(R) \equiv \langle N(R) \rangle^2 - \langle N(R) \rangle^2,$$

where angular brackets denote ensemble average. For Poisson distributions of points (ideal gas), and for traditional correlated disordered systems such as liquids and glasses, the number variance for large $R$ scales as the window volume, i.e., $\sigma_N^2(R) \sim R^d$. A hyperuniform point pattern is one wherein the number variance $\sigma_N^2(R)$ grows more slowly than the window volume for large $R$, i.e., more slowly than $R^d$ [1]. Perfect crystals and quasicrystals are hyperuniform with scaling $\sigma_N^2(R) \sim R^{d-1}$; i.e, the variance grows as the window surface area). In disordered systems however, when the variance scales more slowly than $R^d$, then the system should be considered exotic. Hyperuniform states of disordered matter have hidden order on long length scales that is not apparent at short length scales.

The hyperuniformity condition described above is equivalent to the following condition on the structure factor, $S(\mathbf{k})$, at wave vector $\mathbf{k}$ [1]:

$$\lim_{|\mathbf{k}| \to 0} S(\mathbf{k}) = 0. \quad (2)$$

This condition that there is no scattering at $k = 0$ implies that the infinite-wavelength density fluctuations of the system (when appropriately scaled) vanish. The ensemble-averaged structure factor of a point pattern in $\mathbb{R}^d$ at number density $\rho$ is defined via

$$S(\mathbf{k}) = 1 + \rho h(\mathbf{k}), \quad (3)$$

where $h(\mathbf{k})$ is the Fourier transform of the total correlation function $h(\mathbf{r}) = g_2(\mathbf{r}) - 1$, and $g_2(\mathbf{r})$ is the pair correlation function that characterizes the system. [Note, definition (3) implies that the forward scattering contribution is omitted.] For computational purposes, the structure factor, $S(\mathbf{k})$, for a finite-size point configuration in a fundamental cell under periodic boundary conditions, can be obtained directly from the particle positions $\mathbf{r}_j$ [2]:

$$S(\mathbf{k}) = \frac{1}{N} \left| \sum_{j=1}^{N} \exp(i \mathbf{k} \cdot \mathbf{r}_j) \right|^2 \quad (\mathbf{k} \neq 0). \quad (4)$$

Here $N$ is the total number points in the fundamental cell. Note that to ascertain accurately the very small wave number behavior of $S(k)$ (crucial for the hyperuniformity test), $N$ (system size) must be large.

B. Multiphase media

The hyperuniformity concept has been extended to statistically homogeneous systems made of 2 (or more) phases in which each phase $i$ ($i = 1, 2$) occupies volume fraction $\phi_i$. In the specific case of polymeric particles suspended in water, phase 1 is the polymeric particles while phase 2 is the surrounding aqueous phase. The following is a general formulation of hyperuniformity; it can be applied whether the system is composed of particles or not. In two-phase systems (e.g., composites, binary suspensions, block copolymers, etc.) [18–21], hyperuniformity is manifested in the suppression of local volume-fraction fluctuations:

$$\sigma_\tau^2(R) = \langle \tau(\mathbf{r})^2 \rangle - \langle \tau(\mathbf{r}) \rangle^2. \quad (5)$$

Here the random variable $\tau$ is the local volume fraction of either phase in some observation window of radius $R$ [2], and hence the average of $\tau$ is simply the volume fraction of the phase of interest (either $\phi_1$ or $\phi_2$). Nonhyperuniform disordered media, such as typical liquids and glasses, have the scaling $\sigma_\tau^2(R) \sim R^{-d}$. By contrast, volume-fraction fluctuations in hyperuniform two-phase systems will decrease faster than $R^{-d}$ [2].

It has been shown that this hyperuniformity condition for two-phase media is equivalent to the following condition on the spectral density $\tilde{\chi}(\mathbf{k})$ (defined below) [2]:

$$\lim_{|\mathbf{k}| \to 0} \tilde{\chi}(\mathbf{k}) = 0. \quad (6)$$

Here the spectral density, $\tilde{\chi}(\mathbf{k})$, is the Fourier transform of the autocovariance function

$$\chi(\mathbf{r}) = S_2^{(i)}(\mathbf{r}) - \phi_i^2 \quad (7)$$

where

$$S_2^{(i)}(\mathbf{r}) = \langle I^{(i)}(\mathbf{x}) I^{(i)}(\mathbf{x} + \mathbf{r}) \rangle \quad (8)$$

is the autocorrelation function for phase $i$ [18] and $I^{(i)}(\mathbf{x})$ is the indicator function for phase $i$ defined by

$$I^{(i)}(\mathbf{x}) = \begin{cases} 1, & \mathbf{x} \in \text{phase } i, \\ 0, & \text{otherwise}. \end{cases} \quad (9)$$

where angular brackets denotes an ensemble average.

A two-phase system is one in which the space is partitioned into two disjoint regions (phases); the phase $i$ can be 1 or 2. For a statistically homogeneous two-phase medium, the volume fraction of phase $i$ is simply $\phi_i = \langle I^{(i)}(\mathbf{x}) \rangle$, independent of the position $\mathbf{x}$. The autocorrelation function $S_2^{(i)}(\mathbf{r})$ is essentially the probability that the end points of a vector $\mathbf{r}$ lie in the same phase $i$ when the vector is randomly placed in the medium [18]. The spectral density, $\tilde{\chi}(\mathbf{k})$, is independent of $i$ and can be obtained directly from scattering experiments [22]. For computational purposes, the spectral density, $\tilde{\chi}(\mathbf{k})$, i.e., for a given finite two-phase configuration in a fundamental cell of volume $V$ under periodic boundary conditions, can be obtained directly from the square of the Fourier transform of the configuration [23]. In particular,

$$\tilde{\chi}(\mathbf{k}) = \frac{1}{V} |\tilde{J}^{(i)}(\mathbf{k})|^2, \quad (10)$$

where $\tilde{J}^{(i)}(\mathbf{k})$ is the Fourier transform of $J(\mathbf{x}) = I^{(i)}(\mathbf{x}) - \phi_i$. 

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C. Detecting hyperuniformity in polydisperse sphere packings

Determination of whether a monodisperse packing of spheres is hyperuniform can be established by analyzing the centers of the spheres using the point configuration formulation [Eq. (1) or Eq. (2)] or the space occupied by the spheres using the volume-fraction formulation [Eq. (5) or Eq. (6)]. Both methodologies are equivalent. Importantly, if one desires to ascertain whether a sphere packing with polydisperse size distribution is hyperuniform, then one cannot use the point configuration formulation. In the polydisperse case, the point configuration formulation may lead to a false conclusion that the packing is not hyperuniform due to size disparity. The points in polydisperse samples should be weighted according to the volume occupied by the spheres. In such instances, it has been shown that hyperuniformity can be inferred only from the degree to which local volume-fraction or area-fraction fluctuations are suppressed, as described in Ref. [24].

For polydisperse systems, two local-volume fraction detection methods can be employed. One method is performed in direct space; it consists of measuring the area fraction fluctuations of the sample. The other method is performed in reciprocal space and consists of measuring the spectral density of the sample. Consider a maximally random jammed (MRJ) packing of binary disks under the “strict” jamming constraint; these configurations were generated using the numerical algorithm described in Ref. [24]. It has been established that MRJ packings of particles of general shape and polydisperse size distribution are disordered and hyperuniform [3–6,17,24]. Thus, MRJ systems provide useful models to test the two hyperuniformity analysis methods in the multiphase formulation.

In order to make contact with digital images obtained from our experiments, we digitized the computer-generated jammed binary disk configurations. Therefore, except for digitization, no other uncertainties were introduced. Of course, additional errors will be introduced when an image is extracted from an experiment.

Figure 1 summarizes the real-space and reciprocal-space approaches to ascertain hyperuniformity based on ideal binarized images. A portion of such an image is shown in Fig. 1(a). Particles appear bright and the background appears dark. Figures 1(b) and 1(c) summarize the analysis in real space. A series of circular observation windows of radius $R$ is used to randomly sample the binarized image, which is made dimensionless by the average diameter $D$, and the local area fraction occupied by the disks is recorded for each observation circle. Figure 1(b) shows the probability distribution of the local area fraction $P(\tau)$ as a function of $\tau$ for three different values of $R/D$. In this specific example, $\tau$ fluctuates around a mean value close to the overall packing fraction ($\phi_{MRJ} \approx 0.85$). It can be seen that the distribution gets narrower as $R/D$ increases, thereby demonstrating that the amplitude of the fluctuations of the area fraction decreases for larger observation windows. We calculate the variance $\sigma^2_R$ for this value of $R$, and then repeat this procedure for different $R$ to obtain $\sigma^2_R(R)$ as a function of $R$. These results are shown in Fig. 1(c).

For a disordered nonhyperuniform packing, generated from a hyperuniform packing in which 10% of randomly chosen large and small particles are swapped, we observe the expected scaling $\sigma^2_R \propto (R/D)^{-2}$ (dashed line). By contrast, the variance of the hyperuniform packings exhibits a steeper decrease. A practical way to visualize hyperuniformity is to plot $(R/D)^2\sigma^2_R$ as a function of $R/D$ [Fig. 1(e)]. For the nonhyperuniform packing, this quantity remains constant. In a disordered jammed hyperuniform packing (i.e., the MRJ system) this quantity decreases with $R/D$ [24]. Specifically, one expects the variance to exhibit the asymptotic scaling

$$
\sigma^2_R(R/D) \sim \left(\frac{D}{R}\right)^3 [c_0 + c_1(\ln(R/D))]
+ O \left(\left(\frac{D}{R}\right)^4\right) \quad (R/D \to \infty),
$$

where $c_0$ and $c_1$ are structure-dependent constants of order unity.

Figures 1(d)–(f) summarize the analysis in reciprocal space. Using reciprocal space relation [Eq. (10)], we obtain direction-dependent spectral densities for both hyperuniform and nonhyperuniform packings, such as the ones shown in Figs. 1(d) and 1(e). The origin, $k = 0$, is the center of the two power spectra, and the intensities vary radially as shown for one particular direction [indicated by the arrow in Fig. 1(d)]. Figure 1(d) is obtained from a hyperuniform packing, and it exhibits a black spot at the center of the image. The ring around the black spot corresponds to nearest-neighbor location, i.e., to wave numbers inversely related to the size of the colloids in the image. Figure 1(e) is the power spectrum of the nonhyperuniform packing. As expected, it exhibits a white spot at the image center, indicating that the power spectrum does not vanish at very low wave vectors. Since the two power spectra are isotropic, they are readily angularly averaged to obtain the radial power spectra $\tilde{\chi}(k)$ in Fig. 1(f), where $k = |k|$ is the wave number. By contrast, one can very clearly see that the power spectrum of the hyperuniform packing is unlike that of the spectrum of the nonhyperuniform packing; the hyperuniform packing spectrum vanishes as $k$ tends to zero.

These two methods will be utilized throughout the remainder of this paper, e.g., with experimental data. It is important to note that the methods and work discussed in this subsection employed ideal binarized images as materials for analysis. Therefore, these materials have no uncertainty in particle positions; furthermore, the present numerical data are not degraded by the discretization step that must be carried out when starting from actual experimental images. For example, when the system under investigation is a packed colloid of thermal micron-sized particles, optical microscopy cannot provide us with images of sufficient resolution to generate accurate binarized images of polydisperse samples. Specifically, optical aberration and scattering of light makes it difficult to precisely resolve the edges of each particle, and local variation of the image intensity on larger length scales prevents application of simple thresholding operations to generate binarized images. For these reasons, it is crucial to develop reliable procedures to reconstruct particle packings from the optical image data. Therefore, we will first explore simulated hyperuniform packings in order to assess the effect of the different reconstruction algorithms for assessment of hyperuniformity; these simulated packings will further enable assessment of the most important...
FIG. 1. (Color online) (a) A portion of digitized (binarized) image of a MRJ 2D packing of 10 000 binary hard disks (white circles) in which the fraction of large disks is 0.75 and the size ratio is 1.32. Dashed circles indicate some sample circular observation windows of radius $R$. (b) Distribution function of the local area fraction of white particles $P(\tau)$ as a function of $\tau$, for different values of $R/D$. The distribution gets narrower as $R/D$ increases. $P(\tau)$ can be fitted by Gaussian curves (solid lines), and the center of these curves is the average packing fraction of our system, \( \langle \tau \rangle = \phi_{\text{MRJ}} \approx 0.85 \). (c) The local-volume fraction variance $(R/D)^2 \sigma^2(\tau)$ as a function of $R$ (normalized by the average diameter $D$) for hyperuniform (thick black line) and nonhyperuniform (thin red line) packings. Inset: $\sigma^2(\tau)$ as a function of $R/D$ for hyperuniform (thick black line) and nonhyperuniform (thin red line) packings. The dashed curve, included as a guide to the eye, behaves like $(R/D)^{-2}$. Typical power spectra for hyperuniform (d) and nonhyperuniform (e) packings. The arrow shows $k$ in one specific direction. (f) The spectral density $\tilde{\chi}/D^2$ versus dimensionless wave number $kD/(2\pi)$ as obtained by angular averaging of (d) and (e).

III. MATERIALS AND METHODS

A. Jammed, disordered packings of PNIPAM particles

The experimental samples were composed of soft poly(N-isopropyl acrylamide) (PNIPAM) microgel particles. The samples were binary particle suspensions with PNIPAM particles of two different radii, $r_1 \approx 0.57 \mu m$ and $r_2 \approx 0.43 \mu m$ at $27^\circ C$. The particles were synthesized by surfactant-free radical emulsion polymerization, as described elsewhere [25]. A quasi-2D packing was prepared by confining the binary mixtures of large and small PNIPAM particles between two cover slips (Fisher Scientific) and then sealing the same edges with optical glue (Norland 63) [26]. Since PNIPAM is a temperature-sensitive polymer, the particle diameter can be controlled by changing temperature. Thus, the effective packing fraction of the sample was tuned...
FIG. 2. (Color online) (a) Example of a raw (original) image showing a jammed packing of small and large PNIPAM spheres, each of which has a small degree of polydispersity from their associated mean. (b) Probability density function of individual particle intensity derived from particle tracking together with Gaussian fits for the two populations. The dashed line indicates the cutoff intensity for separation of small and large particles. (c) Pair correlation function for the small particles (red, thin) and the large particles (black, fat). The mean diameter of each population is inferred from the position of the first peak in the pair correlation function. (d) Snippet of the original image with the detected centers superimposed (red crosses: small particles; blue stars: big particles). (e) Reconstructed image using the PCF method. Note, on the scale of this figure, the reconstructions produced by the other two techniques described in the text, DLS and j-PSR methods, are nearly indistinguishable. Therefore, they are not shown. The j-PSR method is comparatively superior for minimizing interparticle overlaps.

in situ by controlling the sample temperature with an objective heater (BiOptechs). Briefly, the temperature was set to 26 °C so that the packing fraction is near the jamming point. The trajectories of $N \approx 4500$ particles in the field of view were extracted from a total of 300 frames of video at 10 frames/s using standard centroid finding and particle tracking techniques [27]. The analysis of the particles mean-square displacement shows that averaging over 30 s of experiments is sufficient for correct assessment of the mean position of the particles. The time-averaged positions of the particles were used for image reconstruction, and the integrated intensity of each particle was used to divide spheres into two groups: particles with large versus small diameter [Fig. 2(a) and 2(b)].

B. Generation of simulated hyperuniform packings

We numerically generated 2D MRJ packings of 200 to 10 000 hard repulsive particles using the algorithm formulated
in Ref. [24]. This packing protocol is a classical molecular dynamics method in which initially small particles in a unit cell under periodic boundary conditions undergo collisions with one another and grow as a function of time at some expansion rate $\gamma$ until they jam. The expansion rate is chosen to be the fastest time consistent with strict jamming, which leads to MRJ states [3]. Each configuration consisted of a 75:25 binary mixture of large and small disks with size ratio of 1.33.

We studied the effect of polydispersity index (PDI) up to PDI = 5%. We first generated such packings, wherein PDI was defined as the standard deviation of the size distribution divided by the corresponding mean diameter. Then, for each configuration, we started the packing algorithm with random positions of nonoverlapping particles and desired polydispersities but with particles about 30% smaller than their final size. We then allowed the particles to grow at an expansion rate, $\gamma = dD/dt = 10^{-2}D_0$, until a jammed state was reached; here $D$ and $D_0$ are, respectively, the final and initial disk diameter.

IV. PACKING RECONSTRUCTION FROM IMAGES

In order to reconstruct the entire packing within an image with high accuracy, it is necessary to determine the center and diameter of each colloidal particle. We employ two classes of reconstruction algorithms for this challenging task. The first class is based on the assumption that populations are ideal, meaning that each population (large and small) are perfectly monodisperse. Within this class, the radius for each population can be assessed either the sample’s pair correlation function (PCF) or dynamic light scattering (DLS). The second class of reconstruction algorithm is an adaptative algorithm. This algorithm recovers the size of the particles by minimizing the overlaps and the gaps between particles, and we refer to it as the jammed-particle size-recovery method, i.e., the “j-PSR method.”

We start the reconstruction from the optical microscopy images. The microscopy images are similar to the one in Fig. 2(a), which shows a typical packing of the binary PNIPAM colloids. Using the algorithm developed by Crocker and Grier [27], the centers of all the particles are readily identified. In our jammed packings, thermal motion is much smaller than the particle size, and we average over many frames to obtain highly accurate particle positions (centers). Next, an intensity based on the particle brightness is assigned to each detected particle; the histogram of intensity is plotted in Fig. 2(b). The histogram clearly exhibits two well-defined populations. These populations correspond to the small and large particles, and each individual particle is labeled as large or small, correspondingly [Fig. 2(d)]. The crucial next step is to assign a radius to each particle.

A. Reconstruction using idealized populations

After the center of each colloidal particle has been identified and classified according to the population to which it belongs, the radius of each particle must be measured. Two ways to carry out this task are explored here. The PCF method requires computation of the pair correlation function (PCF), $g(r)$, which is proportional to the probability density associated with finding the two particle centers at a distance $r$ from one another [18]. The quantity $2\pi r g(r) dr$ gives the expected number of particle centers in an annulus of differential thickness $dr$ located at a radial distance $r$ from a particle center; $\rho$ is the particle number density. Here we compute $g(r)$ only between pairs of small particles or pairs of large particles. The pair correlation function results are shown in Fig. 2(c); the red curve is the pair correlation function of the small particles in the packing, and the black curve is the pair correlation function of the large colloidal particles. The first peaks in the correlation functions give the average distance between nearest neighbors, and in a densely packed system such as ours, we assume the mean diameter of each species to be identical to this distance.

In the DLS method, we assign particle diameters based on data from dynamic light scattering measurements. In this end, a standard DLS setup is employed to measure the mean particle sizes in diluted suspensions of both species at the experiment temperature. Then, using the particle centers and the resultant mean diameters determined by DLS, the packing is reconstructed. An example is shown in Fig. 2(e). The white pixels that appear in the image correspond to zones wherein particle overlap occurs. Because polydispersity is not accounted for in either the PCF or DLS approach, overlaps between particles and the presence of rattler particles are unavoidable. Further, another possible source of error arises from the fact that the separation between large and small particles from Fig. 2(b) is not perfect. Thus, errors may arise from the fact that some small particles are identified as large particles and vice versa. This class of reconstruction procedure does not account for polydispersity, and thus the same limitations apply to both PCF and DLS methods. Note also, because DLS measures a hydrodynamic radius rather than a structural radius, the mean diameter of soft polymeric particles such as PNIPAM particles is likely to be slightly overestimated.

B. Reconstruction using an adaptive algorithm

Our approach follows the spirit of two methods presented in Ref. [17] and Ref. [28]. In the former paper, an iterative algorithm was developed to measure the size of individual particles in two- and three-dimensional imaging experiments wherein particles were not packed too densely. This approach works best for hard spheres with well-known mean particle size and with positional data measured at different times. For soft particle systems, neighboring particles can be separated by gaps or they can overlap. Therefore, by contrast to hard spheres, the diameter of the deformed sphere is somewhat ill-defined, as is the experimentally determined mean radius of the soft spheres. Such is the case for the suspended hydrogel particles employed in the present experiments.

Moreover, our analysis must rely on a single set of time-averaged data to derive particle positions. A situation that is more akin to the conditions of our experiment has been presented in Ref. [17]. Therein, true particle sizes of densely packed polydisperse hard spheres with a single average diameter were successfully reconstructed by taking into account distances (and radii) of nearest and second-nearest neighbor particles; the only inputs required are the
Let the approximate particle radii before the $n$th step of the iterative algorithm (with two steps per cycle) be denoted by $r_i^{n-1}$ and $r_j^{n-1}$, respectively. In the first step, the gaps between particles are corrected using the following equation:

$$r_i^n = r_i^{n-1} + p_{(+)} \frac{\sum_{j=1,m} (d_{ij} - r_i^{n-1} - r_j^{n-1})}{m}.$$  \hspace{1cm} (12)

Here $j$ runs over $m$ nearest neighbors of particle $i$ for which

$$0 \leq d_{ij} - r_i - r_j \leq \alpha(r_i + r_j).$$  \hspace{1cm} (13)

Equation (12) implies that each particle is to be expanded by a fraction, $p_{(+)}$, of the mean gap between itself and its nearest neighbors. To utilize this procedure, we must define nearest-neighbor particles. A number of definitions can be used, but here we call particles $i$ and $j$ nearest neighbors, if the gap between $i$ and $j$ is not larger than a fraction, $\alpha$, of the sum of their radii $r_i^{n-1} + r_j^{n-1}$. This criterion defines the touching neighbors in a more constrained manner; they are not necessarily the geometric neighbors as determined by a Voronoi construction. In practice we perform this expansion for all particles simultaneously, which, in turn, will lead to some overlaps between neighbors.

Thus, the second step in each cycle consists in shrinking particles (i.e., decreasing particle diameter) to minimize the overlap. For this purpose, we recompute a new radius for each particle using the following equation:

$$r_i^{n+1} = r_i^n - p_{(-)} (r_i^n + r_j^n - d_{ij})_{\max}.$$  \hspace{1cm} (14)

Here $(r_i^n + r_j^n - d_{ij})_{\max}$ is the largest absolute overlap between particle $i$ and any of its neighbors $j$. Note that setting the parameter $p_{(-)} = 0.5$ removes all overlap, but by setting it to be a little smaller than 0.5 we can usually obtain better results, since we then avoid overcorrections due to a few particles with grossly overestimated radii. Thus, every cycle consists of the two consecutive steps described by Eqs. (12) and (14). In practice it has proven useful to include an initial shrinking step before the first cycle, replacing $p_{(-)}$ by $p_{(-ini)} \approx 0.5$ in order to start all cycles from an overlap-free configuration.

A useful way to follow the system evolution as a function of cycle (iteration) number is to calculate a so-called “deviation parameter” $\Delta$. $\Delta$ is the sum of the absolute values of gaps and overlaps between neighbors for the entire packing:

$$\Delta = \sum_i \sum_j |d_{ij} - r_i - r_j|.$$  \hspace{1cm} (15)

Note that the iterative procedure minimizes $\Delta$, but $\Delta$ should remain finite even if the fit is perfect, because not all neighbors actually touch.

We tested the algorithm by applying it to simulated packings of particles generated using the procedures described in Sec. III B. The size distribution of an exemplary packing is shown in Fig. 4(a). In this case, we assume a bipolydisperse packing with a size ratio of $\approx 1:3$ and a polydispersity index of PDI = 0.06 for both large and small particles. The size ratio is similar to the experimental value, and we tested our algorithm for a wide range of polydispersities. The first step is to distinguish between small particles (with mean radius $r_i$) and large particles (with mean radius $r_j$). We then binarize the distribution: particles of radius below a certain threshold value...
FIG. 4. (Color online) (a) Particle size histogram with Gaussian fits of bipolydisperse simulated hard sphere packing with $\phi \approx 0.84$ and PDI $\approx 0.06$ for both sizes. To test the algorithm, we binarize the distribution by setting $r \equiv r_f$ for $r_s \leq r_{cut}$ and $r \equiv r_t$ for $r_s > r_{cut}$. We then run the packing code using the following parameters: $p(+) = p(-) = 0.4$, $p(-,ini) = 0.49$, and $\epsilon = 0.05$. (b) and (c) Evolution of fitted radii minus true radii after 1, 4, 7, and 10 full cycles compared to the initial binarized state for big and small particles, respectively. For the initial state and after 10 cycles, Gaussian fits are shown as well (dashed curves); distributions are about an order of magnitude narrower after 10 cycles compared to the initial PDI. (d) The deviation parameter $\Delta$ (see text) decreases initially but is essentially constant after 10 cycles. (c) Fraction of particles in the most populated bin in the histograms of (b) and (c) (and all other cycles), a measure of “distribution sharpness” vs cycles. The quality of the particle size fit is best after about 10 cycles and remains almost constant after that. Note that noninteger cycle numbers belong to the addition step, the initial (binarized) state is at cycle $-0.5$ and cycle 0 corresponds to the initial shrinking step before the first full cycle.

$r_{cut}$ are considered as small and their radius is $r_s$; particles of radius above $r_{cut}$ are considered as large and their radius is $r_t$. By doing such a radius assignment, we pretend that we can only distinguish large and small particles (like in the microscopy experiment). From this starting point, we try now to recover the true particle sizes [Fig. 4(a)]. Starting from these binarized radii, we then run the algorithm using the following parameters: $p(+) = p(-) = 0.4$, $p(-,ini) = 0.49$, and $\epsilon = 0.05$. At each cycle, we can compare the recovered fitted radii $r_f$ to the true values, $r_t$, given by the simulations. We plot the evolution of probability distribution functions of $r_f - r_t$ after 1, 4, 7, and 10 cycles in direct comparison to the starting probability distribution where particles are just binarized for large and small particles [Figs. 4(b) and 4(c)].

A perfect recovery of the true particle radii would correspond to a delta function at $r_f - r_t = 0$. Evidently, after 10 cycles, we recover the true radii with very high accuracy; the top curves in Figs. 4(b) and 4(c) can be fitted by Gaussian curves much narrower than the binarized curve with polydispersity, PDI $\approx \sigma/r \approx 0.06$. After 10 cycles, $\sigma/r < 0.01$, which is comparable to the results of Kurita et al. [17] obtained after time averaging in the “not too densely packed” bipolydisperse packings [28] and in monopolydisperse dense packings without time averaging [17].

V. HYPERUNIFORMITY FROM RECONSTRUCTIONS
BASED ON SIMULATED PACKINGS

A. Finite sample-size effects

The hyperuniformity concept was developed for infinitely large systems. Of course, when working on real experimental systems, one has to account for the finite sample size. Here we employ MRJ packings of spheres known to be hyperuniform and generated by numerical techniques in order to assess finite sample-size effects. One such hyperuniform simulated packing was generated, and we explore the system comprised within a square box with a side length of $2L$. [Fig. 5(a)]. (Note, the simulated packing assumed periodic boundary conditions.) To detect the hyperuniformity of the packing, we use both the direct-space approach for the local-volume fraction variance, Eq. (5) and spectral (reciprocal-space) method via Eq. (10). Within the simulation box, we randomly place circular windows of radius $R$. Because the entire circular window is constrained to be included in the box, the window centers are located inside a smaller square box with side length of $2l$. The quantities $L$, $l$, and $R$ are chosen such that $l + R \leq L$.

FIG. 5. (Color online) (a) Definitions of the size parameters. The square box has a side of size $2L$, the disk of inspection has radius $R$, and its center can span a box of size $l$. The hyperuniform jammed system in the box was simulated under periodic boundary conditions. (b) $(R/D)^2 \sigma^2_{L}$ as a function of $R/L$ for the box of size $2L$ (open circles), and for the larger system containing nine stitched images with box side length $6L$ (solid squares). The signal decreases as expected for a hyperuniform system. Inset: $(R/D)^2 \sigma^2_{L}$ as a function of $R/L$. The drop-off of signal for $R/L = 0.35$ (open circles) and $R/L = 0.70$ (solid squares) is a finite sample-size effect.
In Fig. 5(b) we plot \((R/L)^2 \sigma^2\) as a function of \(R/L\). As mentioned earlier, for a nonhyperuniform system, \((R/D)^2 \sigma^2\) is constant with \(R/L\) for large \(R\), whereas, for a hyperuniform system, \((R/D)^2 \sigma^2\) exhibits a large decrease as a function of \(R/L\), as described by the logarithmic asymptotic relation in Eq. (11).

The results for the finite system with a side length of \(2L\) correspond to the red open-circles in Fig. 5(b). We see that for \(R/L\) between 0 and 0.35, \((R/D)^2 \sigma^2\) is not constant, and it is close to the \(\ln(R)/R\) law expected for hyperuniform packing of spheres [24]. In the inset of Fig. 5(b), we plot \((R/D)^2 \sigma^2\) as a function of \(R/L\). For a hyperuniform system, this quantity should increase as \(\ln R\). At \(R/L \lesssim 0.35\), the data (derived from simulation) shows a strong and sharp decrease followed by oscillations. To assess the origin of this behavior, we construct a sample that is larger in size. Since the original packing was simulated under boundary periodic conditions, we can construct a larger fictitious system of size \(6L \times 6L\) by simply stitching the original sample nine times. The hyperuniformity of the larger system was similarly determined and is represented by the black solid squares in Fig. 5(b).

Interestingly, we can see that the red and black data points superimpose for \(R/L\) ranging from 0 to 0.35. For larger values of \(R/L\), however, the signal for the stitched system continues to follow a logarithm law until it finally decreases for \(R/L \gtrsim 0.7\).

These observations suggest that the drop-off observed for the red data points above 0.35 is indeed due to the finite size of the box. As the radius increases, the size of the zone that the center of the observation window can span decreases. Therefore the zones for which \(\tau\) is measured become more and more similar. This fact explains qualitatively the observed drop-off in fluctuations. For this reason, we will apply the criterion that \(R/L \lesssim 0.35\) to detect hyperuniformity from measurements.

**B. Comparison of reconstruction methods based on simulated packings**

In this section we examine which of the three reconstruction methods is most effective for hyperuniformity analysis of colloidal packings similar to our experimental data. Importantly, we take the finite sample-size effects discussed in the last subsection into consideration for these comparisons. Specifically, we generate hyperuniform packings made of two populations of hard spheres, i.e., particle configurations with no interparticle overlaps. We study populations of particles which have a polydispersity of 0%, 3%, and 5%, respectively. Starting with perfectly known and well-defined particle configurations, we reconstruct the packing using the techniques described earlier. For the PCF method, the values assigned are those inferred from the pair correlation functions. They are slightly larger than the true mean radius by a factor less than 1%. For the DLS method, the values assigned are the exact mean values found from the simulations.

Interestingly, all reconstructions obtained from these three methods yield very similar overall packing fraction. However, both the PCF and the DLS methods deliver significant overlap between particles, i.e., of order 1% of the overall packing fraction. By contrast, the amount of overlap produced by the j-PSR algorithm depends on the \(p_{(-)}\) parameter in the last subtraction step, and the resulting overlap is usually much smaller than the PCF and DLS methods (e.g., it is zero for \(p_{(-)} = 0.5\)). Furthermore, and importantly, the j-PSR reconstruction reliably reproduces the true polydispersity in the particle distributions.

Next, we determine which (if any) reconstructed samples exhibit signatures of hyperuniformity. Figure 6(a) shows the results in direct space for three simulated packings; in these packings each population of small and large particles have a polydispersity of 0%, 3%, and 5%, respectively. In this figure, \((R/D)^2 \sigma^2\) is plotted as a function of \(R/D\) and \(R/L\). Importantly, with knowledge of the correct size of each individual particle, the simulated packings exhibit very similar trends. In particular, \((R/D)^2 \sigma^2\) does not remain constant with the normalized window radius: the packings generated by the simulation are hyperuniform.

We can readily assess how the packing reconstructions derived via different methodologies (PCF, DLS, j-PSR) affect the diagnosis of hyperuniformity. Figure 6(b) compares the packing obtained by the PCF method for two polydisperse populations of colloids against the original simulated packings. For both cases we can see in Fig. 6(b) that the curves obtained from reconstruction strongly differ from those obtained...
from the simulated packings. This observation suggests that neglecting the polydispersity of each population of colloids, and replacing their diameters by the mean diameters obtained using the pair correlation functions, produces small errors in the packing reconstruction that strongly affect the diagnosis of hyperuniformity. We have found that making such an assumption for any sample with polydispersity higher than 3% results in a false negative diagnostic. The same conclusions are drawn from the results presented in Fig. 6(c), wherein the DLS method is used for reconstruction.

By contrast, Fig. 6(d) shows the same curves for the packing obtained using the j-PSR technique. Though the packing obtained by this technique still introduces errors, the curves obtained are very similar to those obtained from the original simulated packings. The reconstructed packings show clear signatures of hyperuniformity. To conclude, this study has shown that when the degree of polydispersity is of the order of 3% or larger, then any reconstruction technique that assumes monodisperse particles produces a false diagnosis, i.e., a true hyperuniform packing does not appear to be hyperuniform if the particle polydispersity distributions have not been accounted for, even when packings are reconstructed using the true radii (e.g., as is the case for the DLS method).

The same conclusions can be drawn by investigating the image power spectrum in the reciprocal space via Eq. (10). Figure 7(a) shows the power spectra for simulated packings for which both populations have a polydispersity of 0%, 3%, and 5%, respectively. The spectral densities or power spectra \( \tilde{\chi} \) of these images are almost indistinguishable, and indeed they vanish at very low wave numbers. However, the power spectra associated with the reconstructed images using either the PCF or DLS methods do not overlap with the power spectrum obtained from a hyperuniform packing [Fig. 7(b) and 7(c)] at low wave numbers. The spectral densities clearly do not vanish. Thus, for a 3% polydisperse sample, reciprocal-space measurements show that the system is not hyperuniform. Finally, Fig. 7(d) shows that the power spectra obtained from packings reconstructed using the j-PSR technique indeed vanish in the long-wavelength limit, which is consistent with the results obtained from direct space measurements.

Overall, the direct-space methods appear to be superior to the reciprocal-space method for detection of hyperuniformity when the system is indeed hyperuniform. Some theoretical arguments shed light on this question. It was observed in Ref. [1] that the window size need only be about an order of magnitude larger than the nearest-neighbor spacing \( D \) in a point pattern to estimate the long-range scaling of the number variance \( \sigma_2^2(R) \) [Eq. (1)] with \( R \). This is due to the fact that the corrections to the leading order hyperuniform term in the asymptotic expansion for large \( R \) are comparatively negligible to this dominant contribution for such moderately sized windows [1]. On the other hand, in order to assess hyperuniformity through the spectral condition (2), one must have very large systems to determine accurately the long-wavelength limit, i.e., wave numbers tending to zero. The sizes of the relatively small windows in the direct-space procedure can be appreciably smaller than the size of the large system needed to get accurate results for the structure factor \( S(k) \) in reciprocal space.

An analogous analysis applies to the local-volume fraction variance \( \sigma_1^2(R) \). To see this explicitly, consider the asymptotic relation (11) that applies to 2D disordered hyperuniform MRJ packings [24]. Recall that the leading-order nonhyperuniform term of order \( (D/R)^2 \) is identically zero or, in practice, very close to zero. (If this were not the case, then the scaling would be dominated by this leading order term and hyperuniformity would not even be a consideration.) Since the constants \( c_0 \) and \( c_1 \) in Eq. (11) are of order unity, and because the implied constant multiplying the correction of order \( (D/R)^4 \) is also of order unity, it is clear than when \( R/D \) is a relatively small number that the hyperuniformity terms will dominate. This effect is to be contrasted with the spectral condition (6), which requires very large systems to access the small-wave number behavior accurately.

To demonstrate these theoretical arguments more explicitly, we analyze the effect of system size with simulated packings consisting of different numbers of bidisperse particles, as shown Fig. 8, and compare the direct-space and reciprocal-space measurements. The curves for both direct-space and reciprocal-space measurements are averaged from 100, 20, 10, and 10 independent configurations of 200-, 500-, 1000-, and 10 000-particle packings, respectively. We average over configurations to remove any variations due to configurational fluctuations, thereby allowing us to more accurately assess system-size effects.

For the direct-space measurements, shown in Fig. 8(a), despite limited system size, even curves for 200- and 500-particle packings are fitted well with Eq. (11), leading to a conclusion of hyperuniform configuration. The extrapolated
FIG. 8. (Color online) (a) Results for the hyperuniformity test performed in the direct space for simulated packings consisting of 200 (black triangle), 500 (red circle), 1000 (blue square), and 10 000 (green diamond) particles. Dashed lines are fitted curves with Eq. (11). (b) Results for the hyperuniformity tests performed by computing the power spectrum of the reconstructed images for the same simulated packings. (c) Magnified view for 200- and 10 000-particle packings in small $kD/2\pi$ region. Red dashed lines for the fit to the first four data points for each system size. Intersections of linear fits with the vertical axis, $\tilde{\chi}/D^2(k=0)$, for all four system sizes are plotted in the insert in (b).

VI. HYPERUNIFORMITY IN EXPERIMENTAL JAMMED PACKINGS OF SOFT SPHERES

Guided by the lessons learned from simulated packings, we finally test the three reconstruction procedures on an experimental jammed (and disordered) packing of PNIPAM particles. The samples were made using procedures described above. The packings are reconstructed using the three methods, DLS, PCF, and j-PSR, and then the hyperuniformity of the reconstructed packings is tested using both direct space and the reciprocal space approaches, as describe in Sec. V.

The results are plotted in Fig. 9. In direct space, measurements of the experimental packing confirm observations made on simulated packings. Specifically, the PCF and DLS methods suggest that the PNIPAM packing is not hyperuniform. Using the j-PSR method, however, the polydispersity index is found to be 6% and 3% for small and large particles, respectively. Furthermore, the packing obtained using the j-PSR method, which includes the effect of polydispersity, strongly suggests that the sample is hyperuniform to within experimental error. Notice that the scaled volume-fraction fluctuations deviate from a linear law, as can be seen in Fig. 9(a). If one fits $\sigma^2$ as a function of $(R/D)^\lambda$, $\lambda$ being the adjustable exponent, we find the following values for $\lambda$: for DLS-reconstructed packings, $\lambda = -1.99 \pm 0.04$, for PCF-reconstructed packing.

The different effects of system size on direct-space and reciprocal-space measurements suggest that direct-space measurements are superior for assessment of hyperuniformity in systems of limited size which are often encountered in experiments. Moreover, when large systems cannot be realized, a more accurate determination of hyperuniformity can be achieved by averaging many similar configurations of relatively small size using direct-space measurements.
Different reconstruction methods that can be employed with finding an optimized approach, we have investigated three from typical microscope images of colloidal packings. In general, any approximation in the reconstruction tends to drive $\lambda$ closer to or even larger than $-2$.

Not surprisingly, hyperuniformity is not confirmed by the reciprocal-space measurements. None of the three power spectra obtained from the three methods vanish at very long wavelengths, although the signals for the j-PSR and the PCF methods plateau at a very low wave number close to $2.5 \times 10^{-3}$. These observations are consistent with the simulations studies that clearly suggest the direct-space approach is superior for detection of hyperuniformity (e.g., in samples of finite size). Taken together, the results strongly suggest that the 2D jammed PNIPAM packing is effectively hyperuniform. This fact is consistent with the observation that, theoretically, the softness of particles does not preclude hyperuniformity. Examples of hyperuniformity in such systems include studies of particles with soft long-ranged interactions [10]. Also, our observations and analysis provide a framework to decipher whether these types of many-particle systems are truly hyperuniform or “nearly” hyperuniform or are not hyperuniform [15,16].

VII. CONCLUSIONS

To conclude, we have devised a general methodology that facilitates diagnosis of hyperuniform particle packings from typical microscope images of colloidal packings. In finding an optimized approach, we have investigated three different reconstruction methods that can be employed with experimental microscopy images. Studies with hyperuniform simulated data enabled us to assess the validity of the methods and to understand the complications of particle polydispersity. Importantly, we found that any procedure that neglects polydispersity as small as 3% fails to reconstruct the original packing with an accuracy that is high enough to determine the hyperuniformity of the packing. This observation is consistent with previous work reported by Berthier et al. [6]. In order to properly account for polydispersity, we developed an algorithm, the so-called j-PSR algorithm, which yields reconstructions that are accurate enough to decipher hyperuniformity. In addition, we discovered and explained why direct-space procedures that determine whether a packing is hyperuniform are more accurate than their reciprocal-space counterparts, i.e., the differences are due to the size of the finite system that can be accessed experimentally. Finally, using the results of our analysis we are able to prove that our 2D jammed and disordered systems composed of soft PNIPAM particles are indeed hyperuniform to within experimental error.

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